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# 1/*f* **noise in semiconductors arising from the heterogeneous detrapping process of individual charge carriers**

# **Aleksejus Kononovicius***<sup>∗</sup>* **and Bronislovas Kaulakys**

Institute of Theoretical Physics and Astronomy, Vilnius University, Vilnius, Lithuania E-mail: aleksejus.kononovicius@[tf](#page-0-0)ai.vu.lt

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**Abstract.** We propose a model of  $1/f$  noise in semiconductors based on the drift of in[dividual charge carriers and their in](https://stacks.iop.org/JSTAT/2024/113201)teraction with the trapping cen[ters. We assume that the trapping centers a](https://doi.org/10.1088/1742-5468/ad890b)re homogeneously distributed in the material. The trapping centers are assumed to be heterogeneous and have unique detrapping rates. We show that uniform detrapping rate distribution emerges as a natural consequence of the vacant trap depths following the Boltzmann distribution, and the detrapping process obeying Arrhenius law. When these laws apply, and if the trapping rate is low in comparison to the maximum detrapping rate,  $1/f$  noise in the form of Hooge's relation is recovered. Hooge's parameter,  $\alpha_{\rm H}$ , is shown to be a ratio between the characteristic trapping rate and the maximum detrapping rate. The proposed model implies that  $1/f$  noise arises from the temporal charge carrier number fluctuations, not from the spatial mobility fluctuations.

**Keywords:** current fluctuations, fluctuation phenomena, renewal processes, numerical simulations

\*Author to whom any correspondence should be addressed.

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# **Contents**



# **1. [Introduction](#page-14-0)**

<span id="page-1-0"></span>The nature of the  $1/f$  noise (often also referred to as low frequency, flicker or pink noise), characterized by power spectral density of  $S(f) \sim 1/f^{\beta}$  form (with  $0.5 \le \beta \le$ 1.5), remains open to discussion despite almost 100 years since the first reports [1–4]. While many materials, devices, and systems exhibit different kinds of fluctuations or noise  $[4-6]$ , only the white noise and the Brownian noise are well understood from the first principles. White noise is characterized by absence of any temporal correlations, and has a flat power spectral dens[it](#page-14-2)y of  $S(f) \sim 1/f^0$  form. Examples of the [wh](#page-14-1)ite noise include thermal and shot noise. Thermal noise is known to arise from the random motio[n](#page-14-2) [of](#page-14-3) the charge carriers. It occurs at any finite temperature regardless of whether the current flows. Shot noise, on the other hand, is a result of the discrete nature of the charge carriers and the Poisson statistics of waiting times before each individual detection of the charge carrier. The Brownian noise is a temporal integral of the white noise, and thus exhibits no correlations between the increments of the signal, it is characterized by a power spectral density of  $S(f) \sim 1/f^2$  form.

Theory of  $1/f$  noise based on the first principles is still an open problem.  $1/f$  noise is of particular interest as it is observed across various physical  $[7-12]$ , and non-physical [13–16] systems. 1/*f* noise cannot be obtained by the simple procedure of integration, differentiation, or simple transformations of well-understood processes. Also the general mechanism of generating 1/*f* noise has not yet been properly id[en](#page-14-4)[tifi](#page-14-5)ed, and there is no [gen](#page-14-6)erally accepted solution to the 1/*f* noise problem.

[The](#page-14-7) oldest explanation for 1/*f* noise involves the superposition of Lorentzian spectra [17–20]. Lorentzian spectral densities themselves may arise from the random telegraph signals [4], and from the Brownian motion with a broad distribution of relaxations [21]. These approaches, as well as many others, are often limited to the specific systems being modeled, or require quite restrictive assumptions to be satisfied [22]. In the recent [dec](#page-14-8)[ade](#page-14-9)s, [s](#page-14-2)eries of models for the  $1/f$  noise based on the specific, autoregressive AR(1),

point process [21], and the agent-based model [23, 24], yielding nonlinear stochastic differential equation [25] was proposed (see [26] for a recent review). Another more recent trend relies on scaling properties and nonlinear transformations of signals [27– 30]. These models, on the other hand, prove to be rather more abstract, and therefore more similar t[o th](#page-14-10)e long-range memory models [fou](#page-14-11)n[d i](#page-14-12)n the mathematical literature, such as fractional Bro[wn](#page-14-13)ian motion [31, 32] or [AR](#page-14-14)CH models [33, 34]. These and other similar models of 1/*f* noise are hardly applicable to the description and explanatio[n o](#page-14-15)f [the](#page-15-0) mostly observable 1/*f* noise in the semiconductors.

On the other hand, for a homogeneous semiconductor material Hooge proposed an empirical relation for the  $1/f$  no[ise](#page-15-1) [dep](#page-15-2)endence on the pa[ram](#page-15-3)[ete](#page-15-4)rs of the material [35, 36],

<span id="page-2-0"></span>
$$
S(f) = \bar{I}^2 \frac{\alpha_{\rm H}}{N f}.\tag{1}
$$

Where  $\overline{I}$  stands for the average current flowing through the cross-section of the semiconductor material, N is the number of charge carriers, and  $\alpha_H$  is the titular Hooge parameter. If the current is kept constant, or does not exhibit large fluctuations, Hooge's empirical relation could be rewritten in terms of voltage or resistivity noise, i.e.  $S_V(f) = \bar{V}^2 \frac{\alpha_H}{Nf}$  or  $S_R(f) = \bar{R}^2 \frac{\alpha_H}{Nf}$  (here the subscripts emphasize fluctuations of which quantity are being observed). However, we are specifically interested in the case of constant voltage, focusing on the power spectral density of the current fluctuations that are associated with equation (1). There were numerous attempts to derive or explain the structure of the Hooge's relation [37–41]. A more recent derivation of the Hooge's parameter, based on the Poisson generation-recombination process modulated by random telegraph noise, was conducted in [42, 43]. Yet these models, as well as many others, cannot be directly applie[d](#page-2-0) to de[scr](#page-15-5)[ibe](#page-15-6) and explain the widespread 1/*f* noise in the semiconductors.

Here, we propose a model of  $1/f$  noise in semiconductors containing heterogeneous trapping centers. As far as the squ[are](#page-15-7) [of t](#page-15-8)he average current  $\overline{I}^2$  is proportional to the squared number of the charge carriers  $N^2$ , Hooge's relation implies that the intensity of 1/*f* noise is proportional to the number of charge carriers *N*. Therefore, as the first approximation we can consider the noise originating from the flow of individual charge carriers. It is known that the drift, and the diffusion, of the charge carriers does not yield 1/*f* noise [4]. Therefore, we consider the drift of the charge carriers interrupted by their entrapment in the trapping centers. We show that, if the detrapping rates of individual trapping centers are heterogeneous and uniformly distributed, 1/*f* noise arises. As an explanation for the uniform detrapping rate distribution, we note that it may arise from the i[nt](#page-14-2)erplay between the Boltzmann distribution of the vacant trap depths (as is observed in various materials [44–47]) and the Arrhenius law (which is often applied in empirical works studying varied activation and detrapping processes in semiconductors [48–51]). In this model, the signal generated by a single charge carrier is similar to the signal composed of non-overlapping rectangular pulses [52]. Here, we derive Hooge's relation, and show that Hooge'[s p](#page-15-9)[ara](#page-15-10)meter is a ratio between the characteristic trapping rate and the maximum detrapping rate. The proportionality between Hooge's parameter [and](#page-15-11) [th](#page-15-12)e characteristic trapping rate was reported earlier [in](#page-15-13) quite a few experimental

works [53–55]. This result prompts us to suggest that 1/*f* noise in semiconductors arises from the fluctuations in the effective number of charge carriers, not from the spatial fluctuations in mobility.

This paper is organized as follows. In section 2 we introduce a model for 1/*f* noise in the se[mic](#page-15-14)[ond](#page-15-15)uctors based on the trapping-detrapping process of a single charge carrier. In section 3 we address the implications of finite experiments and simulations. Namely, we show that the power spectral density produced by a single charge carrier may exhibit spurious low-frequency cutoff. This cutoff disapp[ea](#page-3-1)rs, if the current generated by a large number of charge carriers is considered. Finally, Hooge's empirical relation and Hooge's parameter [v](#page-7-0)alue for the proposed model is derived in section 4. The main results of the paper are summarized in section 5.

# **2. Model for 1**/**f noise in a homogeneous semiconductor [m](#page-11-1)aterial**

<span id="page-3-1"></span><span id="page-3-0"></span>Let us consider a drift of a single charge carrier (e.g. election) through a homogeneous semiconductor material. While the charge carrier is freely moving through the conduction band, it will generate a non-zero contribution to the net current, i.e.  $I_1(t) = a$  for *t* when the charge carrier is free. As the material contains trapping centers, the freely moving charge carrier will eventually get trapped in one of such trapping centers. Let  $\tau_i$  stand for *i*th detrapping time (time spent in the trap) and  $\theta_i$  be *i*th trapping time (time spent moving). Under these considerations the contribution of single charge carrier to the net current will be composed of gaps (duration corresponds to the respective detrapping time) and pulses (duration corresponds to the respective trapping time). For visual illustration of the single charge carrier trapping-detrapping process and a sample signal see figure 1.

The power spectral density of a signal with rectangular pulses of fixed height is given by  $|52|$ 

<span id="page-3-2"></span>
$$
S_1(f) = \lim_{T \to \infty} \left\langle \frac{2}{T} \left| \int_0^T I_1(t) e^{-2\pi i f t} dt \right| \right\rangle = \frac{a^2 \bar{\nu}}{\pi^2 f^2} \text{Re} \left[ \frac{(1 - \chi_\theta(f))(1 - \chi_\tau(f))}{1 - \chi_\theta(f)\chi_\tau(f)} \right]. \tag{2}
$$

In the above *T* stands for observation time (duration of the signal), which is assumed to approach infinity  $[52]$ ,  $\chi_{\tau}(f)$  and  $\chi_{\theta}(f)$  stand for the characteristic functions of the respective detrapping and trapping time distributions, while  $\bar{\nu}$  is the mean number of pulses per unit time. For the ergodic processes, and given a long observation time *T*, the value of  $\bar{\nu}$  is trivially derived from the mean trapping and detrapping times, i.e.  $\bar{\nu} = \frac{1}{\sqrt{\theta_1 + \mu_2}}$  $\frac{1}{\langle \theta \rangle + \langle \tau \rangle}$ . For the n[one](#page-15-13)rgodic processes, or if the observation time *T* is comparatively short, the expected value of  $\bar{\nu}$  can be derived from the means of the appropriately truncated distributions, or it may be defined purely empirically, i.e.  $\bar{\nu} = K/T$  (here *K* is the number of observed pulses).

Typically when trapping–detrapping processes are considered [4, 9] it is assumed that both  $\tau_i$  and  $\theta_i$  are sampled from the exponential distributions with rates  $\gamma_\tau$  and  $\gamma_{\theta}$ , respectively. Characteristic function of the exponential distribution with an event





**Figure 1.** Visualization of the single charge carrier trapping-detrapping process (left) and a sample single charge carrier contribution to the net current (right). Relevant notation:  $\tau_i$  is the detrapping time (gap duration),  $\theta_i$  is the trapping time (pulse duration), *a* is the height of the pulses (single free charge carrier contribution to the net current),  $t_i$  is the time of *i*th detrapping event.

rate  $\gamma$ , is given by

<span id="page-4-0"></span>
$$
\chi(f) = \int_0^\infty \gamma e^{2\pi i f \tau - \gamma \tau} d\tau = \frac{\gamma}{\gamma - 2\pi i f}.
$$
\n(3)

Inserting equation (3) as the characteristic function for both trapping and detrapping time distributions into equation (2) yields a Lorentzian power spectral density [4]. Notably, there were prior works which have examined the case when  $\tau_i$ ,  $\theta_i$ , or both are sampled from distributions with power-law tails [42, 43, 52, 56–59]. Under the power-law distri[bu](#page-4-0)tion assumption, [it](#page-3-2) was shown  $S(f) \sim 1/f^{\beta}$  dependence can be [re](#page-14-2)covered.

Here, let us assume that the trapping centers are het[ero](#page-15-7)g[ene](#page-15-8)o[us.](#page-15-13) [Ea](#page-15-16)[ch](#page-15-17) of them has their own unique depth, or detrapping (activation) energy,  $E_a^{(i)}$ . As is commonly observed [48–51], let us assume that the detrapping process obeys Arrhenius law

$$
\gamma_{\tau}^{(i)} = A \exp\left[-\frac{E_a^{(i)}}{k_B \Theta}\right].\tag{4}
$$

To obtain the overall detrapping time distribution we first need to establish the distribution of detrapping energies. Not all trapping centers will participate in the trappingdetrapping process at all times. Because charge carrier first needs to be trapped, before being detrapped, only vacant trapping centers will participate in the process. In experimental literature [44–47] it is well established that vacant trap level depths (their activation energies) reasonably well follow the Boltzmann distribution

$$
p\left[E_a^{(i)}\right] = C_N \exp\left[-\frac{E_a^{(i)}}{k \cdot B} \right].
$$
\n(5)

In the above  $C<sub>N</sub>$  stands for the normalization constant. Notably, this result also follows directly from the Fermi–Dirac statistics under the assumption that trap level degeneracy is constant in respect to activation energy. Then, from the conservation of the probability

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density, it follows that the distribution of detrapping rates would be uniform

$$
p\left[\gamma_{\tau}^{(i)}\right] = \frac{p\left[E_a^{(i)}\right]}{\left|\frac{\mathrm{d}c}{\mathrm{d}E_a^{(i)}}\right|} = \frac{C_N \exp\left[-\frac{E_a^{(i)}}{k_B \Theta}\right]}{\frac{A}{k_B \Theta} \exp\left[-\frac{E_a^{(i)}}{k_B \Theta}\right]} = \text{const.}\tag{6}
$$

It is important to note that other physical mechanisms could also imply uniform distribution of the detrapping rates as long as  $\frac{p(\eta)}{\left|\frac{d\eta^{(i)}_T}{d\eta}\right|}$  $rac{\mathrm{d} \gamma^{(i)}_{\tau}}{\mathrm{d} \eta}$  $\overline{\mathcal{L}}$  $=$  const (here  $\eta$  is some generic physical

quantity which would impact the detrapping process).

Let  $\gamma_{\tau}^{(i)}$  be uniformly distributed in  $[\gamma_{\min}, \gamma_{\max}]$ . Then it can be shown that the probability density function of the detrapping time distribution is given by

<span id="page-5-0"></span>
$$
p(\tau) = \frac{1}{\gamma_{\max} - \gamma_{\min}} \int_{\gamma_{\min}}^{\gamma_{\max}} \gamma_{\tau} \exp(-\gamma_{\tau}\tau) d\gamma_{\tau}
$$
  
= 
$$
\frac{(1 + \gamma_{\min}\tau) \exp(-\gamma_{\min}\tau) - (1 + \gamma_{\max}\tau) \exp(-\gamma_{\max}\tau)}{(\gamma_{\max} - \gamma_{\min}) \tau^2}.
$$
 (7)

This probability density function saturates for the short detrapping times,  $\tau \ll \frac{1}{\gamma_{\text{max}}}$ . For the longer detrapping times,  $\tau \gg \frac{1}{\gamma_{\min}}$ , it decays as an exponential function. In the intermediate value range,  $\frac{1}{\gamma_{\text{max}}} \ll \tau \ll \frac{1}{\gamma_{\text{min}}}$ , this probability density function has the  $\tau^{-2}$ asymptotic behavior, which is already known to lead to 1/*f* noise [52, 56–58]. The benefit of this formulation is that it allows to see how the  $\tau^{-2}$  asymptotic behavior can emerge in homogeneous semiconductors. Experimentally *τ <sup>−</sup>*<sup>2</sup> asymptotic behavior is observable in quantum dots, nanocrystal, nanorod, and other semiconductors [60–63], with the detrapping times ranging from picoseconds to several months. [Th](#page-15-13)[e as](#page-15-16)[ym](#page-15-18)ptotic behavior of equation (7) can be examined in figure 2 where it is represented by a red curve. Figure 2 also highlights contributions of some of the individual trapping [cen](#page-15-19)[ters](#page-15-20), detrapping time distributions of which are plotted as dashed black curves.

Unlike the simple power-law distribution, this detrapping time distribution does not require the introductio[n](#page-5-0) of any arbitrary cutoffs. Als[o t](#page-6-0)he parameters of this detrapping time distribut[io](#page-6-0)n have explicit physical meaning. Furthermore, the statistical moments are well-defined and have compact analytical forms. The mean of the distribution is given by

<span id="page-5-1"></span>
$$
\langle \tau \rangle = \frac{1}{\gamma_{\text{max}} - \gamma_{\text{min}}} \ln \left( \frac{\gamma_{\text{max}}}{\gamma_{\text{min}}} \right). \tag{8}
$$

Higher order moments also exist and can be easily derived.

The characteristic function of the detrapping time distribution can be obtained either by calculating Fourier transform of equation (7), or by averaging over the characteristic functions of the exponential distribution, equation (3). Both approaches lead to the

<span id="page-6-0"></span>

**Figure 2.** Probability density function of the detrapping time distribution under the assumption that detrapping rates of individual trapping centers are uniformly distributed (red curve), equation (7). The probability density function was calculated for  $\gamma_{\text{min}} = 10^{-3}$ , and  $\gamma_{\text{max}} = 10$  case. Black dashed curves correspond to the exponential probability density functions of the detrapping times from the individual trapping centers with fixed rates:  $\gamma_{\tau} = 10^{-3}$ ,  $2.78 \times 10^{-3}$ ,  $7.74 \times 10^{-3}$ , 2.15 *×* 10*−*<sup>2</sup> , 5.99 *×* 10*−*<sup>2</sup> , 1.67 *×* 10*−*<sup>1</sup> [,](#page-5-0) 4.64 *×* 10*−*<sup>1</sup> , 1.29, 3.59, and 10. Normalization of the exponential probability density functions was adjusted for the visualization purposes, but it remains proportional to their respective contributions.

same expression, but the latter approach is quicker

$$
\chi_{\tau}(f) = \frac{1}{\gamma_{\max} - \gamma_{\min}} \int_{\gamma_{\min}}^{\gamma_{\max}} \frac{\gamma_{\tau}}{\gamma_{\tau} - 2\pi i f} d\gamma_{\tau} = 1 + \frac{2\pi i f}{\gamma_{\max} - \gamma_{\min}} \ln\left(\frac{\gamma_{\max} - 2\pi i f}{\gamma_{\min} - 2\pi i f}\right).
$$
 (9)

If the interval of the possible detrapping rates is broad  $\gamma_{\rm min} \ll \gamma_{\rm max}$ , then for  $\gamma_{\rm min} \ll$  $2\pi f \ll \gamma_{\text{max}}$  the characteristic function can be approximated by

$$
\chi_{\tau}(f) \approx 1 + \frac{2\pi \mathrm{i} f}{\gamma_{\text{max}}} \ln \left( 1 + \frac{\mathrm{i}\gamma_{\text{max}}}{2\pi f} \right) \approx 1 - \frac{2\pi f}{\gamma_{\text{max}}} \left[ \frac{\pi}{2} - \mathrm{i} \ln \left( \frac{2\pi f}{\gamma_{\text{max}}} \right) \right]. \tag{10}
$$

Inserting equation  $(10)$  into equation  $(2)$  we have

<span id="page-6-1"></span>
$$
S_1(f) = \frac{2a^2\bar{\nu}}{\pi\gamma_{\text{max}}f} \operatorname{Re}\left[\frac{\left(1-\chi_{\theta}\left(f\right)\right)\left[\frac{\pi}{2}-i\ln\left(\frac{2\pi f}{\gamma_{\text{max}}}\right)\right]}{1-\chi_{\theta}\left(f\right)\left\{1-\frac{2\pi f}{\gamma_{\text{max}}}\left[\frac{\pi}{2}-i\ln\left(\frac{2\pi f}{\gamma_{\text{max}}}\right)\right]\right\}}\right].
$$
(11)

Assuming that  $\frac{2\pi f}{\gamma_{\text{max}}} [\frac{\pi}{2} - i \ln(\frac{2\pi f}{\gamma_{\text{max}}})] \ll 1$ , which is supported by an earlier assumption that  $2\pi f \ll \gamma_{\text{max}}$ , allows to simplify the above to

<span id="page-6-2"></span>
$$
S_1(f) \approx \frac{a^2 \bar{\nu}}{\gamma_{\text{max}} f}.\tag{12}
$$

This approximation should hold well for  $\gamma_{\rm min} \ll 2\pi f \ll \gamma_{\rm max}$ , and should not depend on the explicit form of  $\chi_{\theta}(f)$  unless  $\chi_{\theta}(f) \approx 1$  for at least some of the frequencies in the range.

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<span id="page-7-2"></span>

**Figure 3.** Power spectral density of the simulated signal (red curve) and its analytical approximation by equation (12) (black dashed curve). Simulated power spectral density was obtained by averaging over  $10<sup>2</sup>$  realizations. Simulation parameters:  $T = 10^6$ ,  $\gamma_{\text{min}} = 10^{-4}$ ,  $\gamma_{\text{max}} = 10^4$ ,  $a = 1$ ,  $\gamma_{\theta} = 1$ .

Let us examine a specific case when the trapping centers are uniformly distributed within the material, and therefore the trapping process can be assumed to be a homogeneous Poisson process. Inserting the characteristic function of the exponential distribution, equation (3), as the characteristic function of the trapping time distribution into equation (2) yields

$$
S_1(f) = \frac{4a^2\bar{\nu}}{\gamma_\theta^2} \operatorname{Re}\left[\frac{1}{1 - \chi_\tau(f) - \frac{2\pi\mathrm{i}f}{\gamma_\theta}}\right].\tag{13}
$$

Then inserting the characteristic function of the proposed detrapping time distribution, equation  $(10)$ , into equation  $(13)$  yields

<span id="page-7-1"></span>
$$
S_1(f) = \frac{a^2 \bar{\nu} \gamma_{\text{max}}}{\gamma_{\theta}^2 f} \times \frac{1}{\left(\frac{\pi}{2}\right)^2 + \left[\frac{\gamma_{\text{max}}}{\gamma_{\theta}} + \ln\left(\frac{2\pi f}{\gamma_{\text{max}}}\right)\right]^2}.
$$
\n(14)

<span id="page-7-0"></span>If the maximum detrapping rate is large in comparison to the trapping rate, i.e.  $\frac{\gamma_{\text{max}}}{\gamma_{\theta}} \gg \frac{\pi}{2}$ and  $\frac{\gamma_{\text{max}}}{\gamma_{\theta}} \gg -\ln(\frac{2\pi f}{\gamma_{\text{max}}})$ , then we recover equation (12). In figure 3 the power spectral density of a simulated signal with comparatively large detrapping rates is shown as a red curve. We have chosen observation time *T* to allow us to show three regimes of the power spectral density: white noise cutoff for  $2\pi f \ll \gamma_{\rm min}$  $2\pi f \ll \gamma_{\rm min}$  $2\pi f \ll \gamma_{\rm min}$ ,  $1/f$  noise for  $\gamma_{\rm min} \ll 2\pi f \ll \gamma_{\rm max}$  and Brown noise for *γ*max *≪* 2*π f*. Longer or similar observation times would yield similar power spectral density.

#### **3. Low-frequency cutoff in finite experiments**

<span id="page-8-0"></span>The obtained approximation, equation  $(12)$ , holds in the infinite observation time limit (single signal of infinite duration  $T$ ) or the infinite number of experiments limit (infinitely many signals with finitely long observation time *T*). If either of the limits does not hold, then the range of frequencies over which the pure  $1/f$  noise is observed becomes narrower. In the finite experiments the [pro](#page-6-2)cess will not reach a steady state, and therefore the cutoff frequencies will depend not on the model parameter values  $\gamma_{\rm min}$  and  $\gamma_{\rm max}$ , but on the smallest and the largest  $\gamma_{\tau}^{(i)}$  values actually observed during the experiment. The difference between  $\gamma_{\text{max}}$  and the largest  $\gamma_{\tau}^{(i)}$  is negligible, because the pure  $1/f$  noise will be observed only if  $γ_{max}$  is a relatively large number. On the other hand the relative difference between  $\gamma_{\text{min}}$  and smallest  $\gamma_{\tau}^{(i)}$  might not be negligible. Let us estimate the expected value of the smallest  $\gamma_{\tau}^{(i)}$  in a finite experiment.

In the model introduced in the previous section  $\gamma_{\tau}^{(i)}$  is sampled from the uniform distribution with  $[\gamma_{\min}, \gamma_{\max}]$  range of possible values. It is known that, for  $x_i$  sampled from the uniform distribution with  $[0, 1]$  range of possible values, the smallest  $x_i$  observed in the sample of size *K* is distributed according to the Beta distribution with the shape parameters  $\alpha_1 = 1$  and  $\alpha_2 = K$  [64]. Thus the expected value of the smallest  $x_i$  is given by

$$
\langle \min \{x_i\}_K \rangle = \frac{\alpha_1}{\alpha_1 + \alpha_2} = \frac{1}{K + 1}.\tag{15}
$$

Rescaling the range of possible values to  $[\gamma_{\min}, \gamma_{\max}]$  yields

$$
\gamma_{\min}^{\text{(eff)}} = \left\langle \min \left\{ \gamma_{\tau}^{(i)} \right\} \right\rangle_K \right\rangle = \frac{\gamma_{\max} - \gamma_{\min}}{K + 1} + \gamma_{\min}.\tag{16}
$$

As *K* corresponds to the number of pulses in the signal, we have that  $K = \bar{\nu}T = \frac{T}{\sqrt{\rho_{\perp}}}$ *⟨θ⟩*+*⟨τ ⟩* and

$$
\gamma_{\min}^{(eff)} = (\gamma_{\max} - \gamma_{\min}) \frac{\langle \theta \rangle + \langle \tau \rangle}{\langle \theta \rangle + \langle \tau \rangle + T} + \gamma_{\min}.
$$
\n(17)

In the above  $\langle \theta \rangle$  is effectively a model parameter as it is trivially given by  $\langle \theta \rangle = \frac{1}{\gamma}$  $\frac{1}{\gamma_\theta}$  , while  $\langle \tau \rangle$  is a derived quantity which has a more complicated dependence on the model parameters  $\gamma_{\min}$  and  $\gamma_{\max}$  (see equation (8)). If the range of possible  $\gamma_{\tau}^{(i)}$  values is broad, i.e.  $\gamma_{\text{max}} \gg \gamma_{\text{min}}$ , we have

$$
\gamma_{\min}^{(eff)} \approx \gamma_{\max} \frac{\gamma_{\max} \langle \theta \rangle + \ln \frac{\gamma_{\max}}{\gamma_{\min}}}{\gamma_{\max} \left( \langle \theta \rangle + T \right) + \ln \frac{\gamma_{\max}}{\gamma_{\min}}} + \gamma_{\min}.
$$
\n(18)

The above applies to the ergodic case with  $\gamma_{\rm min} \gg 1/T$ . In the nonergodic case, for  $\gamma_{\rm min} \lesssim 1/T$ , it would impossible to distinguish between the cases corresponding to the different  $\gamma_{\text{min}}$  values. Therefore, for the nonergodic case,  $\gamma_{\text{min}}$  can be replaced by  $1/T$ yielding



<span id="page-9-0"></span>

**Figure 4.** The effect of increasing the observation time *T* on the obtained power spectral density. Dashed black curve corresponds to equation (12). Simulation parameters:  $a = 1, \gamma_{\theta} = 1, \gamma_{\min} = 0, \gamma_{\max} = 10^3, T = 10^4$  (red curve),  $10^6$  (green curve), and  $10^8$  (blue curve).

$$
\gamma_{\min}^{(\text{eff})} \approx \gamma_{\max} \frac{\gamma_{\max} \langle \theta \rangle + \ln \left( \gamma_{\max} T \right)}{\gamma_{\max} \left( \langle \theta \rangle + T \right) + \ln \left( \gamma_{\max} T \right)} + \frac{1}{T} \approx \frac{1 + \gamma_{\max} \langle \theta \rangle + \ln \left( \gamma_{\max} T \right)}{T} . \tag{19}
$$

For relatively long trapping times,  $\langle \theta \rangle \gg \frac{\ln(\gamma_{\text{max}}T)}{\gamma_{\text{max}}}$ , we have that

$$
\gamma_{\min}^{(\text{eff})} \approx \frac{1 + \gamma_{\max} \langle \theta \rangle}{T} \approx \frac{\gamma_{\max}}{\gamma_{\theta} T}.
$$
\n(20)

From the above, it follows that low-frequency cutoff is always present in singular experiments with one charge carrier, and with finite observation time *T*. The cutoff will be observed at a frequency close to  $\gamma_{\min}^{(\text{eff})}$ . As can be seen in figure 4, the cutoff moves to the lower frequencies as *T* increases, the power spectral density is flat for the lowest observable natural frequencies,  $\frac{1}{T} < f \lesssim \frac{\gamma_{\text{max}}}{\gamma_{\theta} T}$  $\frac{\gamma_{\textrm{max}}}{\gamma_{\theta} T}.$ 

If multiple independent experiments (let *R* be the number of e[xp](#page-9-0)eriments) with finite observation time *T* are performed and the obtained spectral densities are averaged, then the total number of observed pulses increases by a factor of *R* yielding

$$
\gamma_{\min}^{(\text{eff})} = (\gamma_{\max} - \gamma_{\min}) \frac{\langle \theta \rangle + \langle \tau \rangle}{\langle \theta \rangle + \langle \tau \rangle + RT} + \gamma_{\min} \approx \frac{\gamma_{\max} \langle \theta \rangle}{RT} + \frac{1}{T} = \frac{R + \gamma_{\max} \langle \theta \rangle}{RT}.
$$
 (21)

For  $R \gg \gamma_{\text{max}} \langle \theta \rangle$ , no low-frequency cutoff will be noticeable. As shown in figure 5, lowfrequency cutoff disappears as the experiments are repeated and the obtained power spectral densities are averaged.

We have derived equation (12) considering the current generated by a single charge carrier. In many experiments the number of charge carriers *[N](#page-10-0)* will be large,  $N \gg 1$ . Consequently, from the Wiener–Khinchin theorem [4] it follows that performing independent experiments is equivalent to observing independent charge carriers. Therefore for  $N \gg \gamma_{\text{max}} \langle \theta \rangle$  no low-frequ[enc](#page-6-2)y cutoff will be noticeable. Though in this case, the



<span id="page-10-0"></span>

**Figure 5.** The effect of averaging over repeated experiments on the obtained power spectral density:  $R = 1$  (green curve),  $R = 10<sup>3</sup>$  (magenta curve). Dashed black curve corresponds to equation (12). Simulation parameters, with exception to *R*, are the same as for the green curve from figure 4.

power spectral densities of the sign[als](#page-6-2) generated by single charge carriers add up instead of averaging out, yielding a minor generalizatio[n](#page-9-0) of equation (12)

<span id="page-10-1"></span>
$$
S_N(f) \approx \frac{Na^2\bar{\nu}}{\gamma_{\text{max}}f}.\tag{22}
$$

In the above  $\bar{\nu}$  is strictly the mean number of pulses per unit time generated by a single charge carrier.

As can be seen in figure  $6(a)$ , the signal generated by multiple independent charge carriers is no longer composed of non-overlapping pulses, although it retains discrete nature as individual charges drift freely or are trapped by the trapping centers. The amplitude and the slope of the power spectral density are well predicted by equation (22) (as seen in figure  $6(c)$ ). The [di](#page-11-2)stribution of the signal's amplitude would be expected to follow the Binomial distribution with sample size *N* and success probability (probability that the charge carrier is free)

$$
p_{\rm F} = \frac{\langle \theta \rangle}{\langle \theta \rangle + \langle \tau \rangle} \approx 1 - \frac{\langle \tau \rangle}{\langle \theta \rangle}.
$$
\n(23)

The fit by the Binomial distribution shown in figure  $6(b)$  is not perfect, because the nonergodic case is simulated and  $\langle \tau \rangle$  is ill-defined, but predicts the overall shape of the probability distribution rather well. For  $\gamma_{\rm min} \gg 1/T$  the fit would be much better. Notably, with larger *N* and under noisy observation, the Binomial distribution predicted by the model will quickly become indistinguishable [fr](#page-11-2)om the Gaussian distribution. While in some cases  $1/f$  noise is known to behave as a non-Gaussian process, most often it is found to exhibit Gaussian fluctuations [4, 65, 66]. The duration of the reported simulation was chosen arbitrarily, based on the technical considerations. Specifically, we have opted to make  $2^{26}$  observations of the process with sampling period of  $\Delta t = 10^{-4}$ .

<span id="page-11-2"></span>

**Figure 6.** Results of a single simulation with large number of charge carriers *N* and finite duration  $T$ : excerpt of a signal generated by  $10<sup>3</sup>$  independent charge carriers (a), the probability mass function of the amplitude of the signal (b), and the power spectral density of the signal (c). Red curves represent results of numerical simulation, while dashed black curves provide theoretical fits: (b) binomial probability mass function with  $p_F \approx 0.984$  and  $N = 10^3$ , (c) the power spectral density approximation equation (22). Simulation parameters:  $R = 1, N = 10<sup>3</sup>$ ,  $T = 2^{26} \cdot 10^{-4} = 6710.8864, \ a = 1, \ \gamma_{\theta} = 1, \ \gamma_{\min} = 0, \ \gamma_{\max} = 10^{3}.$ 

Notably, [67] also discusses a spurious low[-fr](#page-10-1)equency cutoff that could be observed in single particle experiments. Of the 1/*f* noise models considered in [67] superimposed random telegraph signals and blinking quantum dot models are the most comparable to the model presented here. In [67] each of the superimposed random telegraph signals was assumed [to](#page-16-0) be characterized by their own Poissonian switching rate  $\gamma = \gamma_{\theta} = \gamma_{\tau}$ between the 'on' and 'off' states. It was shown that the condition[al](#page-16-0) power spectral density (requiring a certain min[imu](#page-16-0)m number of pulses,  $K_{\min}$ , to be observed) exhibits low-frequency cutoff at  $f_c \sim K_{\min}/T$ . In our simulations, we typically observe a large number of pulses,  $K \approx \gamma_{\theta}T$ , and should therefore observe the cutoff at  $f_c \sim \gamma_{\theta}$ , but instead, we observe that the cutoff frequency scales as  $1/\gamma_{\theta}$ . The nature of the cutoff is different in the model introduced here. The other, blinking quantum dot, model does not predict low-frequency cutoff, only the ageing effect, which for the pure 1/*f* noise will not be noticeable [52].

#### **4. Derivation of Hooge's empirical relation and Hooge's parameter**

<span id="page-11-1"></span><span id="page-11-0"></span>It is straightforward to see that we can rewrite equation (22) in the form of Hooge's empirical relation, equation  $(1)$ , if we define Hooge's parameter as

<span id="page-11-3"></span>
$$
\alpha_{\rm H} = \frac{N^2 a^2 \bar{\nu}}{\gamma_{\rm max} \bar{I}^2}.
$$
\n
$$
(24)
$$

Further we show that the straightforward expression above can be simplified, and given a more compact form.

As the height of the pulses *a* corresponds to the current generated by a single charge carrier, we have

$$
a = \frac{qv_c}{L},\tag{25}
$$

where  $q$  stands for the charge held by the carrier,  $v_c$  is the free drift velocity between the trappings (which will be much smaller than the thermal velocity of the charge carriers), and *L* is the length of the material. Expression for *a* can be rewritten in terms of the average current flowing through the cross-section of the material  $\sigma_M$ 

 $\overline{I} = \sigma_M n q v_d,$ (26)

where *n* stands for the density of the charge carriers (i.e.  $n = \frac{N}{Lg}$  $\frac{N}{L\sigma_M}$ ), and  $v_d$  is the average drift velocity of the charge carriers. The average drift velocity is related to the free drift velocity via the fraction of time the charge carrier spends drifting

$$
v_{\rm d} = \frac{\langle \theta \rangle}{\langle \theta \rangle + \langle \tau \rangle} v_{\rm c} = \bar{\nu} \langle \theta \rangle v_{\rm c}.
$$
 (27)

Consequently we have

$$
a = \frac{\bar{I}}{N\bar{\nu}\langle\theta\rangle}.\tag{28}
$$

Inserting equation (28) into equation (24) yields the expression of the Hooge's parameter in terms of the characteristic trapping rate and the maximum detrapping rate, assuming that the trapping times are comparatively long  $\langle \theta \rangle \gg \langle \tau \rangle$ ,

<span id="page-12-0"></span>
$$
\alpha_{\rm H} = \frac{1}{\bar{\nu} \left\langle \theta \right\rangle^2 \gamma_{\rm max}} \approx \frac{\gamma_{\theta}}{\gamma_{\rm max}} = \frac{\langle \tau_{\rm min} \rangle}{\langle \theta \rangle}.
$$
 (29)

In the above  $\langle \tau_{\rm min} \rangle = \frac{1}{\gamma_{\rm min}}$  $\frac{1}{\gamma_{\text{max}}}$  is the expected detrapping time generated when a charge carrier is trapped by the shallowest trapping center. The purer materials (i.e. ones with lower trapping center density  $n_c$ ) will have lower  $\alpha_H$  values, as the trapping rate is given by  $\gamma_{\theta} = \langle \sigma_{\rm c} v_{\rm t} \rangle n_{\rm c}$  (here  $v_{\rm t}$  is the thermal velocity of the charge carriers, and  $\sigma_{\rm c}$  is the trapping cross-section). The proportionality  $\alpha_H \propto \gamma_\theta$  was previously reported in [53–55], providing experimental support to equation (29).

Consequently the approximations for the power spectral density generated by the proposed model, equations (12) and (22), can be rewritten in the same form as [Hoo](#page-15-14)[ge'](#page-15-15)s empirical relation. Inserting equation (29) in[to e](#page-12-0)quation (1) yields

<span id="page-12-1"></span>
$$
S_N(f) = \bar{I}^2 \frac{\gamma_\theta}{\gamma_{\text{max}} N f}.\tag{30}
$$

This expression appears to imply tha[t th](#page-12-0)e process unde[r](#page-2-0) consideration is stationary, but this is not true as the average current  $I$  is proportional to the number of pulses per unit time  $\bar{\nu}$ , which in the  $\gamma_{\rm min} \rightarrow 0$  limit is a function of the observation time *T* [52]. Although, for the case of pure 1/*f* noise, the dependence on *T* is logarithmically slow, and barely noticeable. Nevertheless, even if the process would be non-stationary, this should not have any impact on the estimate of Hooge's parameter as only  $\bar{I}$  is impa[cted](#page-15-13) by the non-stationarity.

# **5. Conclusions**

<span id="page-13-0"></span>We have proposed a general model of  $1/f$  noise in homogeneous semiconductors which is based on the trapping-detrapping process of individual charge carriers. In contrast to the many previous works, we have assumed that the detrapping rate of each trapping center is random. We have shown that, if detrapping process obeys Arrhenius law (which is well-established empirically [48–51]), and if the vacant trap depths follow Boltzman distribution (which is also supported by experimental works  $[44-47]$ ), the detrapping rate distribution will be uniform. When detrapping rates are uniformly distributed, a power-law distribution of the detrapping times equation (7) is obtained. It arises from the superposition of exponenti[al d](#page-15-11)[etr](#page-15-12)apping time distributions representing contributions of the individual trapping centers with their own fixed detr[app](#page-15-9)[ing](#page-15-10) rates (see figure 2).

Consequently, regardless of the exact details of the trapping proce[ss](#page-5-0), as long as the trapping process is slow in comparison to the detrapping process, pure  $1/f$  noise in a form of Hooge's empirical relation is obtained, equation (30). Corresponding expression of the Hoo[ge](#page-6-0)'s parameter,  $\alpha_{\rm H}$ , is then found to be a ratio between the rate parameters of the trapping and the detrapping processes, equation (29). The proportionality between the Hooge's parameter and the trapping rate was reported in previous experimental works [53–55], thus providing partial experimen[tal](#page-12-1) verification for the Hooge's parameter expression we have derived from general theoretical considerations. Inverse proportionality between the Hooge's parameter and the maxi[mum](#page-12-0) detrapping rate suggests interesting implications for approaching suppression of 1/*f* noise problem [68–70]. When the Arrh[eni](#page-15-14)[us l](#page-15-15)aw applies, maximum detrapping rate could be increased either by manipulating the pre-exponential factor, or by decreasing shalowest trap depth (minimum activation energy) from which the Boltzmann distribution applies to the trap depth distribution. The obtained expression for the Hooge's parameter also s[ugg](#page-16-1)[est](#page-16-2)s that  $1/f$  noise arises from the temporal charge carrier number fluctuations, not from the spatial mobility fluctuations.

In section 3, we have discussed the implications of finite experiments. We have shown that the power spectral density may exhibit spurious low-frequency cutoff simply due to finite duration of the experiment or simulation. The obtained width of the cutoff is of the same o[rd](#page-7-0)er of magnitude as  $\frac{\gamma_{\text{max}}}{\gamma_{\theta}}$ . This cutoff disappears when the power spectral density is averaged over a large number of experiments, or when the experiment involves a large number of independent charge carriers. In the latter case the distribution of the signal's amplitude follows Binomial distribution, which under imperfect observation will quickly become indistinguishable from the Gaussian distribution.

#### **Data availability statement**

All of the code used to perform the reported numerical simulations is available at https://github.com/akononovicius/flicker-trap-detrap-individual-charge.

# **Author contributions**

**AK:** [Software, Validation, Writing—Original Draft, Writing—Review](https://github.com/akononovicius/flicker-trap-detrap-individual-charge) & Editing, Visualization. **BK:** Conceptualization, Methodology, Writing—Original Draft, Writing—Review & Editing.

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