

Convergence of products of independent random variables to the log-Poisson law

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Let us have a set of independent random values $\{\xi_{nk}\}$ in each sequence, where $n \in \mathbb{N}$, $k = 1, 2, 3, \dots, k_n$ and $k_n \rightarrow \infty$, as $n \rightarrow \infty$. Suppose that the distribution of ξ_{nk} is defined:

ξ_{nk}	$-e$	-1	0	1	e
P	P_{nk}^-	q_{nk}^-	q_{nk}^0	q_{nk}^+	P_{nk}^+

Let the log-Poisson distribution be $\Pi_{\lambda, \mu, \alpha_0, \alpha_1}(u)$ with parameters $\lambda \geq 0$, $\mu \geq 0$, $\lambda + \mu > 0$, $0 < \alpha_0 \leq 1$, $|\alpha_1| \leq \alpha_0 e^{-2\mu}$.

Theorem. *The M-weak limit of random functions $P(\prod_{k=1}^{k_n} \xi_{nk} < u)$ is $\Pi_{\lambda, \mu, \alpha_0, \alpha_1}(u)$ if and only if the next five conditions are satisfied:*

$$\lim_{n \rightarrow \infty} \max_{1 \leq k \leq k_n} \frac{P_{nk}^+ + P_{nk}^-}{1 - q_{nk}^0} = 0, \quad (1)$$

$$\lim_{n \rightarrow \infty} \prod_{k=1}^{k_n} (1 - q_{nk}^0) = \alpha_0, \quad (2)$$

$$\lim_{n \rightarrow \infty} \prod_{k=1}^{k_n} (P_{nk}^+ + q_{nk}^+ - (P_{nk}^- + q_{nk}^-)) = \alpha_1, \quad (3)$$

$$\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} \frac{P_{nk}^+ + P_{nk}^-}{1 - q_{nk}^0} = \lambda + \mu, \quad (4)$$

$$\alpha_1 \left(\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} \frac{q_{nk}^+ P_{nk}^- - q_{nk}^- P_{nk}^+}{(q_{nk}^+ + P_{nk}^-)^2 - (q_{nk}^- + P_{nk}^+)^2} - \mu \right) = 0. \quad (5)$$

Proof. Sufficiency. Let us use Mellin's transformations. Because random variables ξ_{nk} are independent in each sequence, we have:

$$\omega_n^0(t) = \omega_{\prod_{k=1}^{k_n} \xi_{nk}}^0(t) = \prod_{k=1}^{k_n} (q_{nk}^+ + q_{nk}^- + (P_{nk}^+ + P_{nk}^-)e^{it}),$$

$$\omega_n^1(t) = \omega_{\prod_{k=1}^{k_n} \xi_{nk}}^1(t) = \prod_{k=1}^{k_n} (q_{nk}^+ - q_{nk}^- + (p_{nk}^+ - p_{nk}^-)e^{it}),$$

$$W_n(t) = \begin{pmatrix} \omega_n^0(t) & 0 \\ 0 & \omega_n^1(t) \end{pmatrix}.$$

The Mellin’s transformation of the log-Poisson distribution $\Pi_{\lambda, \mu, \alpha_0, \alpha_1}$ is

$$W(t) = \begin{pmatrix} \alpha_0 e^{(\lambda + \mu)(e^{it} - 1)} & 0 \\ 0 & \alpha_1 e^{(\lambda - \mu)(e^{it} - 1)} \end{pmatrix}.$$

Using common theory about M-weak limit of sequence of M-decreasing independent random variables (see, e.g., [2], thm. 6.1), we get that $W_n(t) \rightarrow W(t)$ for every $t \in \mathbb{R}$. Hence

$$P\left(\prod_{k=1}^{k_n} \xi_{nk} < u\right) \xrightarrow{M} \Pi_{\lambda, \mu, \alpha_0, \alpha_1}(u),$$

and the sufficiency is proved.

Necessity. Let we have $P(\prod_{k=1}^{k_n} \xi_{nk} < u) \xrightarrow{M} \Pi_{\lambda, \mu, \alpha_0, \alpha_1}(u)$. Hence $W_n(t) \rightarrow W(t)$. When $t = 0$, we obtain $\omega_n^0(t) \rightarrow \alpha_0, \omega_n^1(t) \rightarrow \alpha_1$.

Conditions (2) and (3) are satisfied. On the other hand:

$$\lim_{n \rightarrow \infty} \prod_{k=1}^{k_n} \left(\frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0} + \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} e^{it} \right) = e^{(\lambda + \mu)(e^{it} - 1)} \tag{6}$$

for all $t \in \mathbb{R}$.

Let $\{\eta_{nk}\}, n \in \mathbb{N}, k = 1, 2, 3, \dots, k_n$ and $k_n \rightarrow \infty$, when $n \rightarrow \infty$ is a set of independent random variables in each sequence. Let us suppose that η_{nk} obtains values 0 or 1 with probabilities $\frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0}$ and $\frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0}$ respectively.

The characteristic function of random variable $\sum_{k=1}^{k_n} \eta_{nk}$ is

$$W_n^*(t) = \prod_{k=1}^{k_n} \left(\frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0} + \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} e^{it} \right).$$

We have $\lim_{n \rightarrow \infty} W_n^*(t) = e^{(\lambda + \mu)(e^{it} - 1)}$ by (6) for all $t \in \mathbb{R}$.

So we can say that

$$\lim_{n \rightarrow \infty} P\left(\sum_{k=1}^{k_n} \eta_{nk} = l\right) = \frac{(\lambda + \mu)^l}{l!} \text{ for all } l = 0, 1, 2, \dots$$

Hence

$$\left\{ \begin{aligned} \lim_{n \rightarrow \infty} P\left(\sum_{k=1}^{k_n} \eta_{nk} = 0\right) &= e^{-(\lambda + \mu)}, \\ \lim_{n \rightarrow \infty} P\left(\sum_{k=1}^{k_n} \eta_{nk} = 1\right) &= (\lambda + \mu)e^{-(\lambda + \mu)}, \\ \lim_{n \rightarrow \infty} P\left(\sum_{k=1}^{k_n} \eta_{nk} = 2\right) &= \frac{(\lambda + \mu)^2}{2!} e^{-(\lambda + \mu)}. \end{aligned} \right. \quad (7)$$

On the other hand,

$$\left\{ \begin{aligned} P\left(\sum_{k=1}^{k_n} \eta_{nk} = 0\right) &= \prod_{k=1}^{k_n} \frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0}, \\ P\left(\sum_{k=1}^{k_n} \eta_{nk} = 1\right) &= \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \prod_{\substack{k'=1 \\ k' \neq k}}^{k_n} \frac{q_{nk'}^+ + q_{nk'}^-}{1 - q_{nk'}^0}, \\ P\left(\sum_{k=1}^{k_n} \eta_{nk} = 2\right) &= \sum_{k=1}^{k_n} \sum_{\substack{k''=1 \\ k'' > k}}^{k_n} \frac{p_{nk'}^+ + p_{nk'}^-}{1 - q_{nk'}^0} \frac{p_{nk''}^+ + p_{nk''}^-}{1 - q_{nk''}^0} \prod_{\substack{k=1 \\ k \neq k', k''}}^{k_n} \frac{q_{nk}^+ + q_{nk}^-}{1 - q_{nk}^0}. \end{aligned} \right. \quad (8)$$

Because

$$\lim_{n \rightarrow \infty} \left(\left(\frac{P(\sum_{k=1}^{k_n} \eta_{nk} = 1)}{P(\sum_{k=1}^{k_n} \eta_{nk} = 0)} \right)^2 - 2 \frac{P(\sum_{k=1}^{k_n} \eta_{nk} = 2)}{P(\sum_{k=1}^{k_n} \eta_{nk} = 0)} \right) = 0,$$

we get

$$\begin{aligned} &\lim_{n \rightarrow \infty} \left(\left(\sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \frac{1 - q_{nk}^0}{q_{nk}^+ + q_{nk}^-} \right)^2 \right. \\ &\quad \left. - \sum_{k'=1}^{k_n} \sum_{\substack{k''=1 \\ k'' > k'}}^{k_n} \frac{p_{nk'}^+ + p_{nk'}^-}{1 - q_{nk'}^0} \frac{p_{nk''}^+ + p_{nk''}^-}{1 - q_{nk''}^0} \frac{1 - q_{nk'}^0}{q_{nk'}^+ + q_{nk'}^-} \frac{1 - q_{nk''}^0}{q_{nk''}^+ + q_{nk''}^-} \right) = 0 \end{aligned}$$

or

$$\sum_{k=1}^{k_n} \left(\frac{p_{nk}^+ + p_{nk}^-}{q_{nk}^+ + q_{nk}^-} \right)^2 = 0.$$

Hence the condition (1) is satisfied.

Let remember connections (7) and (8), then we get:

$$\lim_{n \rightarrow \infty} \frac{P(\sum_{k=1}^{k_n} \eta_{nk} = 1)}{P(\sum_{k=1}^{k_n} \eta_{nk} = 0)} = \lambda + \mu \quad \text{and} \quad \lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{q_{nk}^+ + q_{nk}^-} = \lambda + \mu.$$

Since

$$\sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} - \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{q_{nk}^+ + q_{nk}^-} \leq \max_{1 \leq k \leq k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{q_{nk}^+ + q_{nk}^-},$$

and condition (1) is proved, we get condition (4). It remains to prove condition (5). Because

$$\begin{aligned} \omega_n^1(t) &= \prod_{k=1}^{k_n} ((q_{nk}^+ + p_{nk}^+) - (q_{nk}^- + p_{nk}^-)) \\ &\times \exp \left\{ (e^{it} - 1) \left(\sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} - 2 \sum_{k=1}^{k_n} \frac{q_{nk}^+ p_{nk}^- - q_{nk}^- p_{nk}^+}{(q_{nk}^+ + p_{nk}^+)^2 - (q_{nk}^- + p_{nk}^-)^2} \right) \right. \\ &\left. + B \max_{1 \leq k \leq k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \sum_{k=1}^{k_n} \frac{p_{nk}^+ + p_{nk}^-}{1 - q_{nk}^0} \right\} \rightarrow \alpha_1 e^{(\lambda - \mu)(e^{it} - 1)}, \end{aligned}$$

and using above proved limits (3) and (4), we get

$$\alpha_1 \left(\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} \frac{q_{nk}^+ p_{nk}^- - q_{nk}^- p_{nk}^+}{(q_{nk}^+ + p_{nk}^+)^2 - (q_{nk}^- + p_{nk}^-)^2} - \mu \right) = 0,$$

and condition (5) is satisfied. Theorem is proved.

References

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Nepriklausomų atsitiktinių dydžių sandaugų konvergavimas į logpuasono dėsnį

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Darbe nagrinėjamas serijų sekos nepriklausomų kiekvienoje serijoje atsitiktinių dydžių sandaugų M-silpnas konvergavimas į logpuasono dėsnį. Yra įrodoma teorema su penkiomis sąlygom.